

# Continuous Random Variables

- Differential entropy
- Maximum entropy distributions
- Normal distribution
- Multivariate normal distribution
- Gaussian process
- Differential entropy rate

# Differential Entropy

- Let  $X$  be a continuous random variable with PDF  $f_X(x)$ . The differential entropy of  $X$  is defined as

$$h(X) = E \{-\log_2 f_X\} = - \int_{\mathcal{R}} f_X(x) \log_2 f_X(x) dx$$

- Example: Differential Entropy for uniform distribution

$$f_X(x) = \frac{1}{a} \mathbf{1}_{[0,a]}(x)$$

$$h(X) = \log_2(a)$$

- Note, differential entropy can be negative!

# Properties of Differential Entropy

- Translation does not change the differential entropy

$$h(X + c) = h(X)$$

- Scaling by  $a$  increases the differential entropy by  $\log_2|a|$

$$h(aX) = h(X) + \log_2 |a|$$

# Joint and Conditional Differential Entropy

- Joint differential entropy of  $X$  and  $Y$

$$h(X, Y) = E \left\{ -\log_2 f_{X,Y} \right\}$$

- Conditional differential entropy of  $X$  given  $Y$

$$h(X|Y) = E \left\{ -\log_2 f_{X|Y} \right\}$$

$$h(X|Y) = h(X, Y) - h(Y)$$

- Conditioning reduces differential entropy

$$h(X|Y) \leq h(X)$$

with equality iff  $X$  and  $Y$  are independent.

# Mutual Information

- Mutual information between  $X$  and  $Y$

$$I(X; Y) = E \left\{ \log_2 \frac{f_{X,Y}}{f_X f_Y} \right\}$$

$$I(X; Y) = h(X) - h(X|Y)$$

$$I(X; Y) = h(Y) - h(Y|X)$$

- Mutual information is non-negative

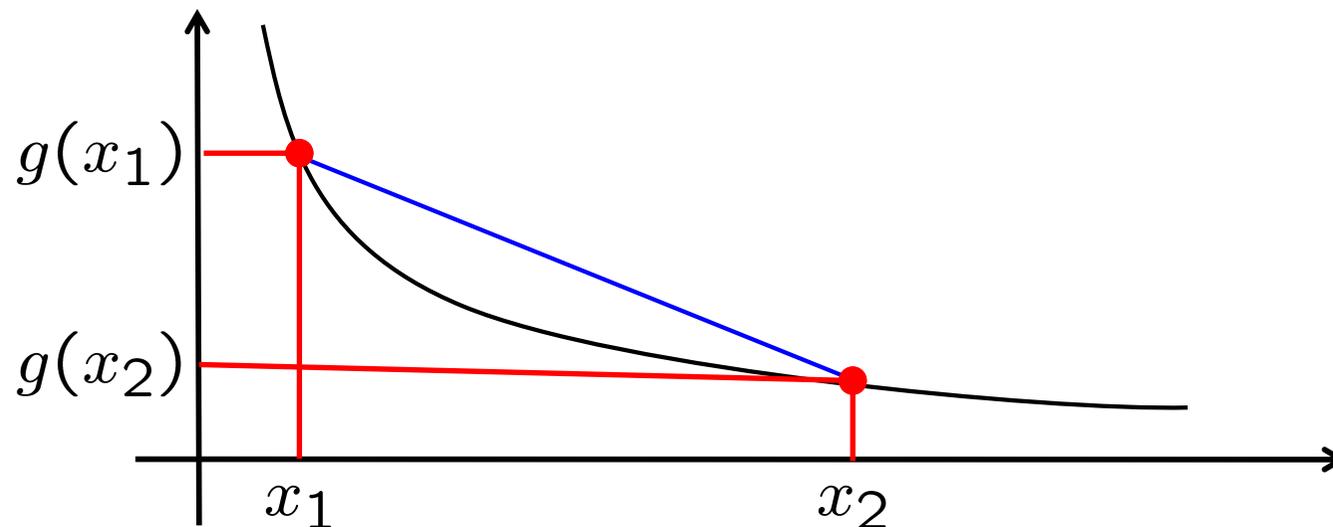
$$I(X; Y) \geq 0$$

with equality iff  $X$  and  $Y$  are independent.

# Remark: Convex Functions

- A function  $g(x)$  is said to be **convex** over an interval  $[a,b]$  if for every  $x_1, x_2 \in [a,b]$  and  $0 \leq \lambda \leq 1$

$$\lambda g(x_1) + [1 - \lambda]g(x_2) \geq g(\lambda x_1 + [1 - \lambda]x_2)$$



- A function is said to be **strictly convex** if equality holds only for  $\lambda = 0$  or  $\lambda = 1$ .

# Remark: Jensen's Inequality

- If  $g$  is a **convex** function and  $X$  is a random variable, then

$$E \{g(X)\} \geq g(E\{X\})$$

- Moreover, if  $g$  is **strictly convex**, then equality implies that  $E\{X\}=X$  with probability 1, i.e.,  $X$  is a constant.

- Example: Two mass point distribution

$$p_1g(x_1) + p_2g(x_2) \geq g(p_1x_1 + p_2x_2)$$

- If  $g$  is strictly convex, equality holds only for the point  $x_1 = x_2 := x$ , that is,  $X$  is a constant.

# Maximum Entropy Distribution for given $\sigma^2$

- For given variance, find PDF that maximizes entropy

$$\inf_{f_X} \{-h(X)\} \quad \text{s.t.} \quad E\{X^2\} = \sigma^2$$

- Solve unconstrained problem  $\inf_{f_X} J(X)$  with  $\lambda > 0$  and

$$\begin{aligned} J(X) &= E \left\{ \ln f_X(X) + \lambda(X^2 - \sigma^2) \right\} \\ &= E \left\{ -\ln \frac{e^{-\lambda(X^2 - \sigma^2)}}{f_X(X)} \right\} \\ &\geq -\ln E \left\{ \frac{e^{-\lambda(X^2 - \sigma^2)}}{f_X(X)} \right\} \\ &= -\ln E\{Y\} \end{aligned}$$

# Maximum Entropy Distribution for given $\sigma^2$

- Jensen: Lower bound is tight iff  $Y$  is a constant.

$$\frac{e^{-\lambda(x^2 - \sigma^2)}}{f_X(x)} = \text{const.} \quad \forall x \in \mathcal{R}$$

- We obtain the PDF

$$f_X(x) = ce^{-\lambda x^2}$$

- **Normal distribution** with variance  $\sigma^2$  that integrates to 1

$$f_X(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{x^2}{2\sigma^2}}$$

# Differential Entropy of Normal Distribution

- For given variance, maximum differential entropy is

$$\begin{aligned}h(X) &= E \{-\log_2 f_X\} \\&= \frac{1}{2} \log_2(2\pi\sigma^2) + \frac{E\{X^2\}}{2\sigma^2} \log_2 e \\&= \frac{1}{2} \log_2(2\pi\sigma^2) + \frac{1}{2} \log_2 e \\&= \frac{1}{2} \log_2(2\pi e\sigma^2)\end{aligned}$$

# Multivariate Normal Distribution

- Let  $\mathbf{X}=(X_1, X_2, \dots, X_n)^T$  have a multivariate normal distribution with zero mean and covariance matrix  $C$

$$f_{\mathbf{X}}(\mathbf{x}) = \frac{1}{(2\pi)^{\frac{n}{2}} |C|^{\frac{1}{2}}} e^{-\frac{1}{2} \mathbf{x}^T C^{-1} \mathbf{x}}$$

- Change of coordinate system:  $\mathbf{Y} = T^T \mathbf{X}$ , where  $T^T C T = \Lambda$  being diagonal with Eigenvalues  $\lambda_i$ ; note that  $|T| = 1$

$$\mathbf{x}^T C^{-1} \mathbf{x} = \mathbf{y}^T T^T C^{-1} T \mathbf{y} = \mathbf{y}^T \Lambda^{-1} \mathbf{y}$$

$$f_{\mathbf{Y}}(\mathbf{y}) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi} \sqrt{\lambda_i}} e^{-\frac{1}{2} \frac{y_i^2}{\lambda_i}}$$

# Multivariate Normal Distribution

- Differential entropy of a multivariate normal random variable

$$\begin{aligned}h(\mathbf{Y}) &= h(T^T \mathbf{X}) \\ &= h(\mathbf{X}) + \log_2 |T^T| \\ &= h(\mathbf{X})\end{aligned}$$

$$\begin{aligned}h(\mathbf{X}) &= \frac{n}{2} \log_2(2\pi e) + \sum_{i=1}^n \frac{1}{2} \log_2 \lambda_i \\ &= \frac{n}{2} \log_2(2\pi e) + \frac{1}{2} \log_2 |C| \\ &= \frac{1}{2} \log_2 [(2\pi e)^n |C|]\end{aligned}$$

# Bivariate Normal Random Variable

- Covariance matrix C

$$C = \begin{bmatrix} \sigma_1^2 & \sigma_1\sigma_2\rho_{12} \\ \sigma_1\sigma_2\rho_{12} & \sigma_2^2 \end{bmatrix} \quad |\rho_{12}| \leq 1$$

- Mutual information between components

$$\begin{aligned} I(X_1; X_2) &= h(X_1) + h(X_2) - h(X_1, X_2) \\ &= \frac{1}{2} \log_2 \frac{\left| \begin{bmatrix} \sigma_1^2 & 0 \\ 0 & \sigma_2^2 \end{bmatrix} \right|}{|C|} \\ &= -\frac{1}{2} \log_2(1 - \rho_{12}^2) \end{aligned}$$

- $I(X_1; X_2) = 0$  for  $\rho_{12} = 0$

# Continuous-Valued Stochastic Process

- Continuous-valued discrete-time stochastic process  $\{X_i\}$ ,  $X_i \in \mathcal{R}$ , is an indexed sequence of continuous random variables  $X_1, X_2, \dots, X_n$  with joint PDF

$$f_{X_1, X_2, \dots, X_n} = \Pr \{ (X_1, X_2, \dots, X_n) = (x_1, x_2, \dots, x_n) \}$$

$$\forall (x_1, x_2, \dots, x_n) \in \mathcal{R}^n$$

- A stochastic process is said to be stationary if the joint PDF is invariant with respect to shifts

$$f_{X_1, X_2, \dots, X_n} = f_{X_{1+l}, X_{2+l}, \dots, X_{n+l}} \quad \forall l \in \mathcal{Z}$$

# Differential Entropy Rate

- How does the differential entropy of  $X^n$  grow with  $n$ ?
- Differential entropy rate of a continuous-valued stochastic process  $\{X_i\}$  is defined by

$$h(\{X_i\}) = \lim_{n \rightarrow \infty} \frac{1}{n} h(X_1, X_2, \dots, X_n)$$

- Using the chain rule

$$h(\{X_i\}) = \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n h(X_i | X_{i-1}, \dots, X_1)$$

# Gaussian Stochastic Process

- Consider a zero-mean cyclostationary Gaussian process  $\{X_i\}$  whose statistical properties repeat with period  $n$

$$\begin{aligned} h(\{X_i\}) &= \frac{1}{n} h(X_1, X_2, \dots, X_n) \\ &= \frac{1}{2} \log_2(2\pi e) + \frac{1}{n} \log_2 |C^{(n)}| \end{aligned}$$

- Let the covariance matrix be **circulant** with  $C_{kl} = E\{X_k X_l\}$

$$C^{(n)} = \begin{bmatrix} c_0 & c_1 & c_2 & \cdots & c_{n-1} \\ c_{n-1} & c_0 & c_1 & \cdots & c_{n-2} \\ c_{n-2} & c_{n-1} & c_0 & \cdots & c_{n-3} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ c_1 & c_2 & c_3 & \cdots & c_0 \end{bmatrix} \quad C_{kl} = c_{(l-k) \bmod n}$$

# Remark: Eigenvalues of Circulant Matrix

- Let  $C$  have Eigenvalues  $\lambda_\nu$  and Eigenvectors  $t^{(\nu)}$

$$Ct^{(\nu)} = \lambda_\nu t^{(\nu)}$$

$$\sum_{l=1}^n C_{kl} t_l = \lambda t_k \quad \forall \quad k = 1, \dots, n$$

$$\sum_{l=0}^{n-k} c_l t_{l+k} + \sum_{l=n-k+1}^{n-1} c_l t_{l+k-n} = \lambda t_k \quad \forall \quad k = 1, \dots, n$$

- Solve difference equation by using complex roots of unity

$$\lambda_\nu = \sum_{l=0}^{n-1} c_l e^{-j\frac{2\pi}{n}l\nu} := \text{DFT}\{c_l\}$$

# Gaussian Stochastic Process

- Consider a zero-mean cyclostationary Gaussian process with period  $n$

$$\begin{aligned}h(\{X_i\}) &= \frac{1}{2} \log_2(2\pi e) + \frac{1}{n} \frac{1}{2} \log_2 |C^{(n)}| \\ &= \frac{1}{2} \log_2(2\pi e) + \frac{1}{n} \sum_{\nu=0}^{n-1} \frac{1}{2} \log_2 \lambda_{\nu} \\ &= \frac{1}{2} \log_2(2\pi e) + \frac{1}{n} \sum_{\nu=0}^{n-1} \frac{1}{2} \log_2 \text{DFT}\{c_l\}_{\nu}\end{aligned}$$

# Remark: Szegő's Theorem

- Let the period  $n$  grow and consider a stationary Gaussian stochastic process with autocorrelation sequence and PSD

$$\phi_{XX}[l] = E\{X_i X_{i+l}\} \quad \Phi_{XX}(\omega) = \sum_{l \in \mathbb{Z}} \phi_{XX}[l] e^{-j\omega l}$$

- Then, the covariance matrix is Toeplitz with Eigenvalues  $\lambda_\nu$

$$C^{(n)} = \begin{bmatrix} \phi_0 & \phi_1 & \phi_2 & \cdots & \phi_{n-1} \\ \phi_{-1} & \phi_0 & \phi_1 & \cdots & \phi_{n-2} \\ \phi_{-2} & \phi_{-1} & \phi_0 & \cdots & \phi_{n-3} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \phi_{1-n} & \phi_{2-n} & \phi_{3-n} & \cdots & \phi_0 \end{bmatrix}$$

- The distribution of discrete Eigenvalues converges

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{\nu=0}^{n-1} g(\lambda_\nu) = \frac{1}{2\pi} \int_{-\pi}^{\pi} g(\Phi(\omega)) d\omega$$

# Gaussian Stochastic Process

- Given a zero-mean stationary Gaussian process  $\{X_i\}$  with

$$\phi_{XX}[l] = E\{X_i X_{i+l}\} \quad \Phi_{XX}(\omega) = \sum_{l \in \mathbb{Z}} \phi_{XX}[l] e^{-j\omega l}$$

- The differential entropy rate is

$$\begin{aligned} h(\{X_i\}) &= \frac{1}{2} \log_2(2\pi e) + \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{\nu=0}^{n-1} \frac{1}{2} \log_2 \lambda_\nu \\ &= \frac{1}{2} \log_2(2\pi e) + \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{\nu=0}^{n-1} \frac{1}{2} \log_2 \text{DFT}\{c_l\}_\nu \\ &= \frac{1}{2} \log_2(2\pi e) + \frac{1}{2\pi} \int_{-\pi}^{\pi} \frac{1}{2} \log_2 \Phi_{XX}(\omega) d\omega \end{aligned}$$