

# $\ell_1$ sparse methods in system identification

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Tutorial on “Sparse and low-rank representation methods in control, estimation and system identification”  
ECC, July 17-19, 2013

# Outline

- Problem statement
- Convex relaxation
- Applications
- How to tune the regularization parameter?
- When / why does it work?
- Some theory
- A model selection tradeoff
- An alternative: Bayesian methods
- Conclusions

# Problem statement

## Assumptions:

- System:  $Y_N = \Phi_N \theta^o + E_N$ 
  - ▶  $E_N \sim (0, \sigma^2 I)$
  - ▶  $n$ : # parameters
  - ▶  $N$ : # samples
  - ▶ *Sparsity*: Most entries of  $\theta^o$  are **zero**
  
- Model:  $Y_N = \Phi_N \theta + E_N$

**Problem:** Estimate zeros of  $\theta$  (detection - model selection)  
& non-zero entries (estimation)

## $\ell_0$ regularization

**Idea:** Impose sparsity as constraint on  $\#$  nonzero entries of  $\theta$ :

$$\begin{aligned} \min_{\theta} V_N(\theta) \\ \text{s.t. } \|\theta\|_0 \leq c \end{aligned}$$

Here:

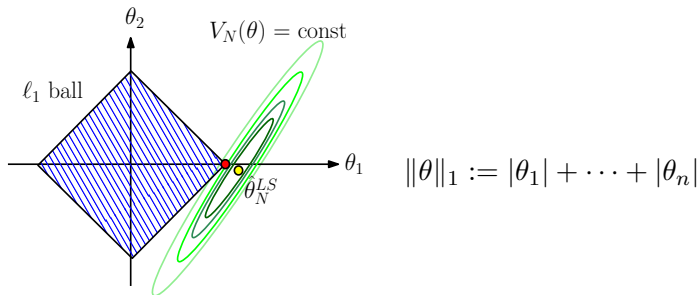
- $V_N(\theta) := \frac{1}{N} \|Y_N - \Phi_N \theta\|_2^2$  (least squares criterion)
- $\|\theta\|_0 := \#$  non-zero parameters

**Problem:** Combinatorial explosion (intractable if  $n$  is large)

# Convex relaxation

Replace hopeless problem with relaxation!

$$\begin{aligned} \min_{\theta} V_N(\theta) \\ \text{s.t. } \|\theta\|_1 \leq \lambda \end{aligned} \quad (\text{LASSO})$$



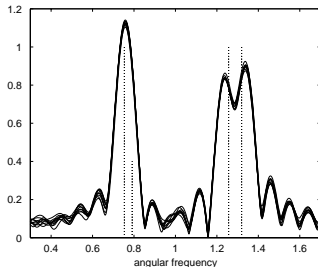
*Still remaining:* How to determine  $\lambda$ ?

# Applications

## 1. Spectral line estimation

$$y_t = \sum_{k=1}^N \alpha_k e^{j\omega_k t} + e_t,$$

$$\alpha_k \in \mathbb{C}$$



*Idea:* Grid the frequency range!

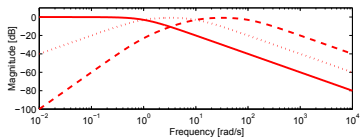
$$Y_N = \begin{bmatrix} y_{t_1} \\ \vdots \\ y_{t_N} \end{bmatrix} \quad \Phi_N = \begin{bmatrix} e^{j\omega_1 t_1} & \dots & e^{j\omega_n t_N} \\ \vdots & \ddots & \vdots \\ e^{j\omega_1 t_N} & \dots & e^{j\omega_n t_N} \end{bmatrix}$$

Impose sparsity constraint on  $\theta = [\alpha_1 \dots \alpha_n]^T$

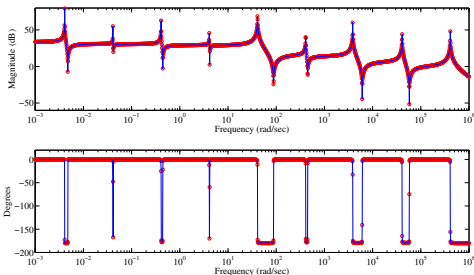
# Applications (cont.)

## 2. Basis function selection / separable least squares

Same idea as for spectral line estimation, but using general basis functions: Laguerre, Kautz, ...



Basis functions



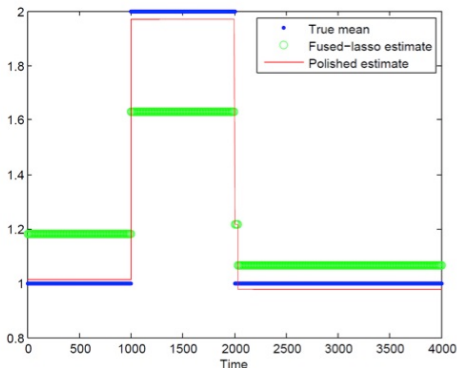
Fit

Ref: Welsh, Rojas, Hjalmarsson & Wahlberg, SYSID, 2012

# Applications (cont.)

## 3. Change detection

$y_t \sim \mathcal{N}(m_t, \sigma^2)$ ,  
where  $m_{t+1} = m_t$  often



*Fused LASSO:*

$$\min_{m_t} \frac{1}{2} \sum_{t=1}^N [y_t - m_t]^2 + \lambda \sum_{t=2}^N |m_t - m_{t-1}|$$



# How to tune $\lambda$ ?

- **AIC / BIC:**

$$\min_{\lambda} V_N(\hat{\theta}_{\lambda}) + \text{pen}(\text{DF}(\hat{\theta}_{\lambda}))$$

where

$$\text{DF}(\hat{\theta}_{\lambda}) = \|\hat{\theta}_{\lambda}\|_0$$

$$\text{pen}(n) = 2n/N \text{ (AIC) or } = n \ln(N)/N \text{ (BIC)}$$

- **Cross-validation:**

$$\min_{\lambda} V_N^{val}(\hat{\theta}_{\lambda})$$

## How to tune $\lambda$ ? (cont.)

- **SPARSEVA:** (for  $n < N$ )

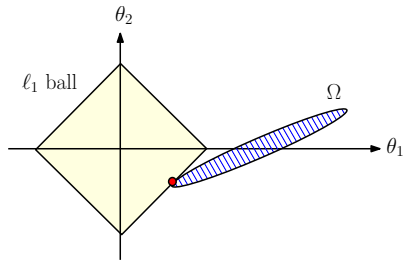
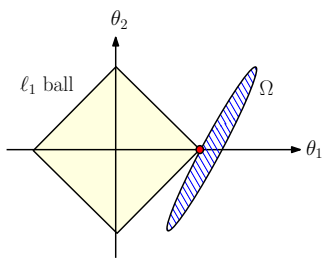
$$\min_{\theta} \|\theta\|_1$$

$$\text{s.t. } V_N(\theta) \leq V_N(\hat{\theta}_N^{LS})(1 + \varepsilon_N)$$

where  $\varepsilon_N = 2n/N$  (AIC) or  $= n \ln(N)/N$  (BIC)

- **Data independent choices:** E.g.  $\lambda \propto N^c$  ( $1/2 < c < 1$ )

# When / why does it work?



Sparse solution NOT obtained

- Shape of level curves of  $V_N$  depend on regressors  $\Phi$

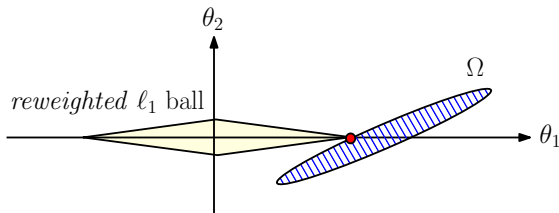
# When / why does it work? (cont.)

One solution: **Adaptive LASSO** (H. Zou, JASA, 2006)

$$\begin{aligned} \min_{\theta} V_N(\theta) \\ \text{s.t. } \sum_k \frac{|\theta_k|}{|\hat{\theta}_k^{LS}|} \leq \lambda \end{aligned}$$

*Interpretations:*

- (1) Resembles “ $\|\theta\|_0 < \lambda$ ” !
- (2) Reweighting of the  $\ell_1$  ball



## Some theory

### Definition (Consistence)

$\hat{\theta}_N$  is consistent in probability if  $\hat{\theta}_N \xrightarrow{p} \theta^o$  as  $N \rightarrow \infty$

Consistence is mostly useful if  $n \ll N$ . Otherwise, the following notion is relevant:

### Definition (Persistence)

$\hat{\theta}_N$  is persistent if  $E\{V_N(\hat{\theta}_N)\} - E\{V_N(\theta_N^*)\} \rightarrow 0$  as  $N \rightarrow \infty$ , where

$$\theta_N^* = \arg \min_{\theta} E\{V_N(\theta)\}$$

- LASSO-type estimators are typically consistent if AIC/BIC is used (for fixed  $n$ ), and persistent when using CV

## Some theory (cont.)

### Definition (Model selection consistence / sparsistence)

$\hat{\theta}_N$  is model selection consistent if  $\mathbf{P}\{\text{supp } \hat{\theta}_N = \text{supp } \theta^o\} \rightarrow 1$   
as  $N \rightarrow \infty$

- Adaptive LASSO with  $\lambda$  chosen via BIC is sparsistent for fixed  $n$ , while it is not with AIC
- For  $n \rightarrow \infty$ , (Adaptive-) LASSO is rarely sparsistent: at most one can enforce  $\text{supp } \hat{\theta}_N \supseteq \text{supp } \theta^o$  in probability

## Some theory (cont.)

### Definition (Oracle property)

$\hat{\theta}_N$  has the oracle property if

$$\sqrt{N}(\hat{\theta}_N - \theta^o) \xrightarrow[N \rightarrow \infty]{d} \mathbf{N}(0, M^\dagger)$$

That is,  $\hat{\theta}_N$  has the same asymptotic distribution as the least-squares oracle, which knows the sparsity pattern of  $\theta^o$

- If (Adaptive-) LASSO is sparsistent, one can achieve the oracle property by *polishing*: The non-zero entries of  $\hat{\theta}_N$  are re-estimated using least squares

# A model selection tradeoff

## Definition (Minimax rate optimality)

$\hat{\theta}_N$  is minimax rate optimal if  $E\{V_N(\hat{\theta}_N)\} - E\{V_N(\theta_N^*)\} \rightarrow 0$  at the fastest possible rate, uniformly in  $\theta_0$

- Minimax rate optimality  $\implies$  Optimal prediction ability
- Model selection consistence  $\implies$  Recovery of 'truth'

*Can we have both?*

**NO!** This is a fundamental limitation in estimation, independent of the estimator (Yang, 2005; Leeb & Ptscher, 2008)



## An alternative: Bayesian methods

**Idea:** Assume that  $\theta$  has a *prior* distribution

$$\theta_i \sim \mathcal{N}(0, \lambda_i), \quad \lambda_i \geq 0, \quad i = 1, \dots, n$$

- $\hat{\lambda}_i$  determined by maximizing  $p(Y_N; \lambda_i)$
- $\hat{\theta}_i$  estimated as  $E\{\theta_i | Y_N, \hat{\lambda}_i\}$
- $\lambda_i = 0 \implies \theta_i = 0!$  (the prior induces sparsity!)
- Seems to induce better sparse estimates than LASSO (i.e., more sparse for same amount of shrinkage), but relies on non-convex programming (local minima!)

Ref: Aravkin, Burke, Chiuso & Pillonetto, CDC, 2011

# Conclusions

- $\ell_1$  regularization as a means to impose sparsity
- Applications to model / regressor / basis function selection  
+ estimation
- How to choose the regularization parameter?
- Theoretical properties and tradeoffs

## Conclusions (cont.)

- Extensions to nonlinearly parameterized models and other kinds of sparsity (piecewise constant signals, graphical models, ...)
- Alternatives: (Empirical-) Bayesian approaches, iterative / greedy methods